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LETTER TO THE EDITOR

Quantum Fisher–Bures information of two-level systems and a three-level extension

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Abstract. Braunstein and Caves have recently demonstrated that the Bures metric on the mixed quantum states is equivalent—up to a proportionality factor of four—to the statistical distinguishability or quantum Fisher information metric. The volume element of these metrics can then—adapting a fundamental Bayesian principle of Jeffreys to the quantum context—serve as a reparametrization-invariant prior measure over the quantum states. The implications of this line of reasoning for the two-level systems, in general, and an embedding of them into a certain set of three-level systems are investigated.

In this letter, we study, among other topics, the four-dimensional convex set of three-level (spin-1) density matrices of the particular form $(0 \le v \le 1)$

$$\rho = \frac{1}{2} \begin{pmatrix} v+z & 0 & x-iy \\ 0 & 2-2v & 0 \\ x+iy & 0 & v-z \end{pmatrix}.$$
 (1)

For v = 1, the middle level is inaccessible and, in effect, the two-level (spin- $\frac{1}{2}$) density matrices

$$\rho = \frac{1}{2} \begin{pmatrix} 1+z & x-iy\\ x+iy & 1-z \end{pmatrix}$$
(2)

are recovered. The domain of admissible (due to trace and non-negativity constraints) values of the parameters x, y and z is then the unit ball (Bloch or Poincaré sphere), $x^2+y^2+z^2 \le 1$. So, equation (1) serves as one of many possible extensions or generalizations of (2) (cf [1]). (Physical photons, although spin-1 particles, are, due to their masslessness, describable by (2).)

Let us attach to the domain of 2×2 density matrices (2), the Bures metric given by [2, formula (3.7)], cf [3,4]

$$\frac{1}{4}\operatorname{Tr}\left\{d\rho\,d\rho + \frac{1}{|\rho|}(d\rho - \rho\,d\rho)(d\rho - \rho\,d\rho)\right\}.$$
(3)

The matrix $(g_{ij}; i, j = x, y, z)$ corresponding to this metric is

$$\frac{1}{4(1-x^2-y^2-z^2)} \begin{pmatrix} 1-y^2-z^2 & xy & xz \\ xy & 1-x^2-z^2 & yz \\ xz & yz & 1-x^2-y^2 \end{pmatrix}.$$
 (4)

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Its inverse (g^{ij}) assumes the simple form

$$4\begin{pmatrix} 1-x^{2} & -xy & -xz\\ -xy & 1-y^{2} & -yz\\ -xz & -yz & 1-z^{2} \end{pmatrix}.$$
(5)

The associated volume element [5] of the Bures metric is obtained by taking the square root of the determinant of (4). The result is

$$1/8(1-x^2-y^2-z^2)^{1/2}\,\mathrm{d}x\,\mathrm{d}y\,\mathrm{d}z\,. \tag{6}$$

Braunstein and Caves [6] have shown that the Bures metric (which extends to mixed states the Fubini–Study metric on pure states) is simply proportional (by a factor of four) to the Fisher information (statistical distinguishability) metric on the quantum states. Relying upon this essential equivalence, together with Jeffreys' principle [7–9] for using the square root of the determinant of the (classical) Fisher information matrix as a reparametrization-invariant prior, we advance (6) as a prior measure over the two-level density matrices (2). From it—through a normalization—one obtains a prior probability distribution

$$p(x, y, z) = 1/\pi^2 (1 - x^2 - y^2 - z^2)^{1/2}$$
(7)

over the unit ball, $x^2 + y^2 + z^2 \le 1$. (The average of the von Neumann entropy, $-\operatorname{Tr} \rho \ln \rho$, over the unit ball is then $2 \ln 2 - 7/6 \approx 0.2196277$, cf [10].)

Since

$$\int_{-(1-x^2-y^2)^{1/2}}^{(1-x^2-y^2)^{1/2}} p(x, y, z) \, \mathrm{d}z = 1/\pi \tag{8}$$

the (three) bivariate marginal probabilities of (7) are uniform distributions over unit discs $(x^2 + y^2 \leq 1, ...)$ —agreeing, in this particular manner, with Laplace's principle of insufficient reason [11]. Then, the three univariate marginal distributions are of the form

$$\int_{-(1-x^2)^{1/2}}^{(1-x^2)^{1/2}} 1/\pi \, \mathrm{d}y = 2(1-x^2)^{1/2}/\pi \qquad (-1 \leqslant x \leqslant 1) \,. \tag{9}$$

Under the transformation, x = 2q - 1, this becomes a beta distribution (figure 1),

$$8q^{1/2}(1-q)^{1/2}/\pi \qquad (0 \le q \le 1) \tag{10}$$

with its two parameters equalling $\frac{3}{2}$. The family of beta distributions is typically employed for the role of prior distributions over binomial (0/1) parameters [9].

As an illustration of the application of Bayes' theorem [8, 9] to the estimation of quantum systems [12–14], let us hypothesize an experimental situation in which spin measurements are performed on each of fourteen replicas of a two-level quantum system—three are taken in the *X*-direction with two 'ups' recorded, five in the *Y*-direction with three 'ups' and six in the *Z*-direction with two 'ups'. Then the posterior (modified) probability distribution over the unit ball is proportional to the product of the prior (7) and the likelihood

$$\left(\frac{1-x}{2}\right)\left(\frac{1+x}{2}\right)^2\left(\frac{1-y}{2}\right)^2\left(\frac{1+y}{2}\right)^3\left(\frac{1-z}{2}\right)^4\left(\frac{1+z}{2}\right)^2\tag{11}$$

since in a two-level system with parameters x, y, z, the probability of an 'up' in the X-direction is (1+x)/2 and a 'down', (1-x)/2,.... This product can be normalized, through an integration over the unit ball, to comprise the posterior probability distribution

$$7168(1-x)(1+x)^{2}(1-y)^{2}(1+y)^{3}(1-z)^{4}(1+z)^{2}/1903\pi^{2}(1-x^{2}-y^{2}-z^{2})^{1/2}.$$
(12)



Figure 1. Univariate marginal probability distribution for two-level systems, expressed as a beta distribution (10).

Let us now attempt to extend this line of analysis to the three-level density matrices of the form (1). The Bures metric for such systems is given by [2, formula (3.8)]

$$\frac{1}{4} \operatorname{Tr} \left\{ d\rho \, d\rho + \frac{3}{1 - \operatorname{Tr} \rho^3} (d\rho - \rho \, d\rho) (d\rho - \rho \, d\rho) + \frac{3|\rho|}{1 - \operatorname{Tr} \rho^3} (d\rho - \rho^{-1} \, d\rho) (d\rho - \rho^{-1} \, d\rho) \right\}.$$
(13)

The result is representable by the matrix $(g_{ij}; i, j = v, x, y, z)$

$$\frac{1}{4(v^{2} - x^{2} - y^{2} - z^{2})} \times \begin{pmatrix} (x^{2} + y^{2} + z^{2} - v)/(1 - v) & -x & -y & -z \\ -x & (y^{2} + z^{2} - v^{2})/v & xy/v & xz/v \\ -y & xy/v & (x^{2} + z^{2} - v^{2})/v & yz/v \\ -z & xz/v & yz/v & (x^{2} + y^{2} - v^{2})/v \end{pmatrix} (14)$$

having the particularly simple inverse $(g^{ij}; i, j = v, x, y, z)$

$$4 \begin{pmatrix} (1-v)v & (1-v)x & (1-v)y & (1-v)z \\ (1-v)x & v-x^2 & -xy & -xz \\ (1-v)y & -xy & v-y^2 & -yz \\ (1-v)z & -xz & -yz & v-z^2 \end{pmatrix}.$$
 (15)

(Removing the factor four from (15) yields the inverse of the quantum Fisher information matrix [6, 15–17]. This then serves—in a non-Bayesian application—as a (Cramér–Rao) lower bound, in the sense of non-negative definiteness, on the covariance matrix of unbiased estimates of the parameters, v, x, y, z [15–17].)

The square root of the determinant of (14) is (cf equation (6))

$$\frac{1}{16v(1-v)^{1/2}(v^2-x^2-y^2-z^2)^{1/2}}.$$
(16)

This can be normalized, using spherical coordinates (r, θ, ϕ) to perform the integrations, to

$$3/4\pi^2 v(1-v)^{1/2} (v^2 - x^2 - y^2 - z^2)^{1/2}.$$
(17)



Figure 3. Univariate marginal (β) probability distribution (19) over auxiliary parameter (v) for extended system.

(The ranges, $0 \le r = (x^2 + y^2 + z^2)^{1/2} \le v$ and $0 \le v \le 1$ were employed.) In spherical coordinates, equation (17) assumes the form

$$3r^{2}\sin\theta/4\pi^{2}v(1-v)^{1/2}(v^{2}-r^{2})^{1/2}.$$
(18)

(Figure 2 shows the marginal distribution of (18) over r and v.) The univariate marginal distribution of (17) and (18) over the variable v is, again (cf equation (10)), an (asymmetric) beta distribution (figure 3),

$$3v(1-v)^{-1/2}/4$$
 $(0 \le v \le 1)$ (19)

with its two parameters equalling 2 and $\frac{1}{2}$. Holding v fixed (V), the conditional distributions of (17) and (18) are

$$1/\pi^2 V^2 (V^2 - x^2 - y^2 - z^2)^{1/2}$$
⁽²⁰⁾

$$r^{2}\sin\theta/\pi^{2}V^{2}(V^{2}-r^{2})^{1/2}.$$
(21)

(For V = 1, equation (20) reduces to (7).) However, integrations could not be exactly nor numerically performed over v to obtain the corresponding marginal distributions over r, θ , ϕ or x, y, z.

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