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## LETTER TO THE EDITOR

# Quantum Fisher-Bures information of two-level systems and a three-level extension 

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#### Abstract

Braunstein and Caves have recently demonstrated that the Bures metric on the mixed quantum states is equivalent-up to a proportionality factor of four-to the statistical distinguishability or quantum Fisher information metric. The volume element of these metrics can then-adapting a fundamental Bayesian principle of Jeffreys to the quantum context—serve as a reparametrization-invariant prior measure over the quantum states. The implications of this line of reasoning for the two-level systems, in general, and an embedding of them into a certain set of three-level systems are investigated.


In this letter, we study, among other topics, the four-dimensional convex set of three-level (spin-1) density matrices of the particular form $(0 \leqslant v \leqslant 1)$

$$
\rho=\frac{1}{2}\left(\begin{array}{ccc}
v+z & 0 & x-\mathrm{i} y  \tag{1}\\
0 & 2-2 v & 0 \\
x+\mathrm{i} y & 0 & v-z
\end{array}\right)
$$

For $v=1$, the middle level is inaccessible and, in effect, the two-level (spin- $\frac{1}{2}$ ) density matrices

$$
\rho=\frac{1}{2}\left(\begin{array}{cc}
1+z & x-\mathrm{i} y  \tag{2}\\
x+\mathrm{i} y & 1-z
\end{array}\right)
$$

are recovered. The domain of admissible (due to trace and non-negativity constraints) values of the parameters $x, y$ and $z$ is then the unit ball (Bloch or Poincaré sphere), $x^{2}+y^{2}+z^{2} \leqslant 1$. So, equation (1) serves as one of many possible extensions or generalizations of (2) (cf [1]). (Physical photons, although spin-1 particles, are, due to their masslessness, describable by (2).)

Let us attach to the domain of $2 \times 2$ density matrices (2), the Bures metric given by [ 2 , formula (3.7)], cf [3, 4]

$$
\begin{equation*}
\frac{1}{4} \operatorname{Tr}\left\{\mathrm{~d} \rho \mathrm{~d} \rho+\frac{1}{|\rho|}(\mathrm{d} \rho-\rho \mathrm{d} \rho)(\mathrm{d} \rho-\rho \mathrm{d} \rho)\right\} \tag{3}
\end{equation*}
$$

The matrix $\left(g_{i j} ; i, j=x, y, z\right)$ corresponding to this metric is

$$
\frac{1}{4\left(1-x^{2}-y^{2}-z^{2}\right)}\left(\begin{array}{ccc}
1-y^{2}-z^{2} & x y & x z  \tag{4}\\
x y & 1-x^{2}-z^{2} & y z \\
x z & y z & 1-x^{2}-y^{2}
\end{array}\right)
$$

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Its inverse $\left(g^{i j}\right)$ assumes the simple form

$$
4\left(\begin{array}{ccc}
1-x^{2} & -x y & -x z  \tag{5}\\
-x y & 1-y^{2} & -y z \\
-x z & -y z & 1-z^{2}
\end{array}\right)
$$

The associated volume element [5] of the Bures metric is obtained by taking the square root of the determinant of (4). The result is

$$
\begin{equation*}
1 / 8\left(1-x^{2}-y^{2}-z^{2}\right)^{1 / 2} \mathrm{~d} x \mathrm{~d} y \mathrm{~d} z \tag{6}
\end{equation*}
$$

Braunstein and Caves [6] have shown that the Bures metric (which extends to mixed states the Fubini-Study metric on pure states) is simply proportional (by a factor of four) to the Fisher information (statistical distinguishability) metric on the quantum states. Relying upon this essential equivalence, together with Jeffreys' principle [7-9] for using the square root of the determinant of the (classical) Fisher information matrix as a reparametrizationinvariant prior, we advance (6) as a prior measure over the two-level density matrices (2). From it-through a normalization-one obtains a prior probability distribution

$$
\begin{equation*}
p(x, y, z)=1 / \pi^{2}\left(1-x^{2}-y^{2}-z^{2}\right)^{1 / 2} \tag{7}
\end{equation*}
$$

over the unit ball, $x^{2}+y^{2}+z^{2} \leqslant 1$. (The average of the von Neumann entropy, $-\operatorname{Tr} \rho \ln \rho$, over the unit ball is then $2 \ln 2-7 / 6 \approx 0.2196277$, cf [10].)

Since

$$
\begin{equation*}
\int_{-\left(1-x^{2}-y^{2}\right)^{1 / 2}}^{\left(1-x^{2}-y^{2}\right)^{1 / 2}} p(x, y, z) \mathrm{d} z=1 / \pi \tag{8}
\end{equation*}
$$

the (three) bivariate marginal probabilities of (7) are uniform distributions over unit discs $\left(x^{2}+y^{2} \leqslant 1, \ldots\right)$-agreeing, in this particular manner, with Laplace's principle of insufficient reason [11]. Then, the three univariate marginal distributions are of the form

$$
\begin{equation*}
\int_{-\left(1-x^{2}\right)^{1 / 2}}^{\left(1-x^{2}\right)^{1 / 2}} 1 / \pi \mathrm{d} y=2\left(1-x^{2}\right)^{1 / 2} / \pi \quad(-1 \leqslant x \leqslant 1) . \tag{9}
\end{equation*}
$$

Under the transformation, $x=2 q-1$, this becomes a beta distribution (figure 1),

$$
\begin{equation*}
8 q^{1 / 2}(1-q)^{1 / 2} / \pi \quad(0 \leqslant q \leqslant 1) \tag{10}
\end{equation*}
$$

with its two parameters equalling $\frac{3}{2}$. The family of beta distributions is typically employed for the role of prior distributions over binomial ( $0 / 1$ ) parameters [9].

As an illustration of the application of Bayes' theorem $[8,9]$ to the estimation of quantum systems [12-14], let us hypothesize an experimental situation in which spin measurements are performed on each of fourteen replicas of a two-level quantum system-three are taken in the $X$-direction with two 'ups' recorded, five in the $Y$-direction with three 'ups' and six in the $Z$-direction with two 'ups'. Then the posterior (modified) probability distribution over the unit ball is proportional to the product of the prior (7) and the likelihood

$$
\begin{equation*}
\left(\frac{1-x}{2}\right)\left(\frac{1+x}{2}\right)^{2}\left(\frac{1-y}{2}\right)^{2}\left(\frac{1+y}{2}\right)^{3}\left(\frac{1-z}{2}\right)^{4}\left(\frac{1+z}{2}\right)^{2} \tag{11}
\end{equation*}
$$

since in a two-level system with parameters $x, y, z$, the probability of an 'up' in the $X$ direction is $(1+x) / 2$ and a 'down', $(1-x) / 2, \ldots$. This product can be normalized, through an integration over the unit ball, to comprise the posterior probability distribution
$7168(1-x)(1+x)^{2}(1-y)^{2}(1+y)^{3}(1-z)^{4}(1+z)^{2} / 1903 \pi^{2}\left(1-x^{2}-y^{2}-z^{2}\right)^{1 / 2}$.


Figure 1. Univariate marginal probability distribution for two-level systems, expressed as a beta distribution (10).

Let us now attempt to extend this line of analysis to the three-level density matrices of the form (1). The Bures metric for such systems is given by [2, formula (3.8)]

$$
\begin{equation*}
\frac{1}{4} \operatorname{Tr}\left\{\mathrm{~d} \rho \mathrm{~d} \rho+\frac{3}{1-\operatorname{Tr} \rho^{3}}(\mathrm{~d} \rho-\rho \mathrm{d} \rho)(\mathrm{d} \rho-\rho \mathrm{d} \rho)+\frac{3|\rho|}{1-\operatorname{Tr} \rho^{3}}\left(\mathrm{~d} \rho-\rho^{-1} \mathrm{~d} \rho\right)\left(\mathrm{d} \rho-\rho^{-1} \mathrm{~d} \rho\right\} .\right. \tag{13}
\end{equation*}
$$

The result is representable by the matrix $\left(g_{i j} ; i, j=v, x, y, z\right)$

$$
\begin{align*}
& \frac{1}{4\left(v^{2}-x^{2}-y^{2}-z^{2}\right)} \\
& \times\left(\begin{array}{cccc}
\left(x^{2}+y^{2}+z^{2}-v\right) /(1-v) & -x & -y & -z \\
-x & \left(y^{2}+z^{2}-v^{2}\right) / v & x y / v & x z / v \\
-y & x y / v & \left(x^{2}+z^{2}-v^{2}\right) / v & y z / v \\
-z & x z / v & y z / v & \left(x^{2}+y^{2}-v^{2}\right) / v
\end{array}\right) \tag{14}
\end{align*}
$$

having the particularly simple inverse $\left(g^{i j} ; i, j=v, x, y, z\right)$

$$
4\left(\begin{array}{cccc}
(1-v) v & (1-v) x & (1-v) y & (1-v) z  \tag{15}\\
(1-v) x & v-x^{2} & -x y & -x z \\
(1-v) y & -x y & v-y^{2} & -y z \\
(1-v) z & -x z & -y z & v-z^{2}
\end{array}\right)
$$

(Removing the factor four from (15) yields the inverse of the quantum Fisher information matrix [6, 15-17]. This then serves-in a non-Bayesian application-as a (Cramér-Rao) lower bound, in the sense of non-negative definiteness, on the covariance matrix of unbiased estimates of the parameters, $v, x, y, z[15-17]$.

The square root of the determinant of (14) is (cf equation (6))

$$
\begin{equation*}
1 / 16 v(1-v)^{1 / 2}\left(v^{2}-x^{2}-y^{2}-z^{2}\right)^{1 / 2} \tag{16}
\end{equation*}
$$

This can be normalized, using spherical coordinates $(r, \theta, \phi)$ to perform the integrations, to

$$
\begin{equation*}
3 / 4 \pi^{2} v(1-v)^{1 / 2}\left(v^{2}-x^{2}-y^{2}-z^{2}\right)^{1 / 2} \tag{17}
\end{equation*}
$$



Figure 2. Bivariate marginal probability distribution over auxiliary parameter ( $v$ ) and radial parameter $(r)$ for the extended system.


Figure 3. Univariate marginal $(\beta)$ probability distribution (19) over auxiliary parameter $(v)$ for extended system.
(The ranges, $0 \leqslant r=\left(x^{2}+y^{2}+z^{2}\right)^{1 / 2} \leqslant v$ and $0 \leqslant v \leqslant 1$ were employed.) In spherical coordinates, equation (17) assumes the form

$$
\begin{equation*}
3 r^{2} \sin \theta / 4 \pi^{2} v(1-v)^{1 / 2}\left(v^{2}-r^{2}\right)^{1 / 2} \tag{18}
\end{equation*}
$$

(Figure 2 shows the marginal distribution of (18) over $r$ and $v$.) The univariate marginal distribution of (17) and (18) over the variable $v$ is, again (cf equation (10)), an (asymmetric) beta distribution (figure 3),

$$
\begin{equation*}
3 v(1-v)^{-1 / 2} / 4 \quad(0 \leqslant v \leqslant 1) \tag{19}
\end{equation*}
$$

with its two parameters equalling 2 and $\frac{1}{2}$. Holding $v$ fixed $(V)$, the conditional distributions of (17) and (18) are

$$
\begin{align*}
& 1 / \pi^{2} V^{2}\left(V^{2}-x^{2}-y^{2}-z^{2}\right)^{1 / 2}  \tag{20}\\
& r^{2} \sin \theta / \pi^{2} V^{2}\left(V^{2}-r^{2}\right)^{1 / 2} \tag{21}
\end{align*}
$$

(For $V=1$, equation (20) reduces to (7).) However, integrations could not be exactly nor numerically performed over $v$ to obtain the corresponding marginal distributions over $r, \theta, \phi$ or $x, y, z$.

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